

Theory of Machine Learning

Exercise sheet 8 — July 5, 2024

Exercise I (On Popoviciu's inequality) The objective of this exercise is to prove Popoviciu's bound on the variance of an almost surely bounded random variable.

Given a square integrable random variable X such that $\mathbb{P}(m \leq X \leq M) = 1$ with $M, m \in \mathbb{R}$, Popoviciu's inequality states:

$$\text{Var}(X) \leq \frac{(M - m)^2}{4}.$$

1. Proof of the inequality using Bhatia–Davis inequality:

- (a) Show that $\mathbb{E}[(M - X)(X - m)] \geq 0$.
- (b) Show that $\mathbb{E}[(M - X)(X - m)] = -\mathbb{E}[X^2] - mM + (m + M)\mathbb{E}[X]$.
- (c) Deduce that $\mathbb{E}[X^2] \leq -mM + (M + m)\mathbb{E}[X]$.
- (d) (Bhatia–Davis inequality) Knowing that $\text{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$, show that:

$$\text{Var}(X) \leq (M - \mathbb{E}[X])(\mathbb{E}[X] - m).$$

- (e) (AM–GM inequality) Given $a, b \in \mathbb{R}^2$, prove that $ab \leq \left(\frac{a+b}{2}\right)^2$.
 - (f) Apply the AM–GM inequality on the variance upper bound given by Bhatia–Davis inequality to get Popoviciu's inequality.
2. (The equality case) Given a discrete random variable X with $\mathbb{P}(X = m) = 1/2$ and $\mathbb{P}(X = M) = 1/2$, show that:

$$\text{Var}(X) = \frac{(M - m)^2}{4}.$$

3. Application of the Popoviciu's inequality to a continuous uniformly distributed random variable $X \sim \mathcal{U}([a, b])$ with $a, b \in \mathbb{R}^2$ (and $a < b$). The density of X is:

$$\forall x \in \mathbb{R}, \quad f(x) := \frac{1}{b - a} \mathbb{1}_{x \in [a, b]}.$$

- (a) Compute the expected value $\mathbb{E}[X] = \int_{\mathbb{R}} xf(x) dx$.
 - (b) Compute the second moment $\mathbb{E}[X^2] = \int_{\mathbb{R}} x^2 f(x) dx$.
 - (c) Knowing that $\text{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$, compute the variance of X .
(Hint: $\text{Var}(X) = \frac{(b-a)^2}{12}$.)
 - (d) Apply Popoviciu's inequality to bound the variance of X . What are your thoughts on the tightness of this bound?
4. Compare this bound to the one seen in the lecture when the loss function is bounded. To remind you:

$$\text{Var}(\ell(Y, f(X))) \leq 4\ell_{\infty}^2,$$

where the loss $\ell(\cdot, \cdot)$ is bounded by $\ell_{\infty} \in \mathbb{R}_+$ and (X, Y) is a data sample.

Exercise II (On the generalization bound of linear regression in fixed design) The objective of this exercise is to prove a generalization bound for an OLS estimator in the fixed design setting.

Consider vector-valued inputs and real-valued outputs ($\mathcal{X} = \mathbb{R}^d$ and $\mathcal{Y} = \mathbb{R}$) with $X := (X_1, \dots, X_n)^\top \in \mathbb{R}^{n \times d}$ the input vector and $Y := (Y_1, \dots, Y_n)^\top \in \mathbb{R}^n$ the response vector. Let $\phi(x) = (x_1, \dots, x_d)^\top$ and $\Phi \in \mathbb{R}^{n \times d}$ the matrix of inputs with row i defined as $\Phi_{i,:} := \phi(X_i)^\top$. We work in the fixed design setting where for a fixed input $X \in \mathbb{R}^{n \times d}$, the output is $Y = \Phi\theta + \varepsilon$ (ε i.i.d. $\mathcal{N}(0, 1)$) and $\theta \in \mathbb{R}^d$. The true risk $\mathcal{R}(\theta)$ of θ and its empirical counterpart $\widehat{\mathcal{R}}(\theta)$ are defined as:

$$\mathcal{R}(\theta) := \mathbb{E}_\varepsilon \left[\frac{1}{n} \|Y - \Phi\theta\|_2^2 \right],$$

$$\widehat{\mathcal{R}}(\theta) := \frac{1}{m} \sum_{i=1}^m \frac{1}{n} \|Y^{(i)} - \Phi\theta\|_2^2,$$

where $\{Y^{(i)}\}_{i=1}^m$ are i.i.d. random vectors sampled as $Y^{(i)} = \Phi\theta + \varepsilon^{(i)}$ ($\varepsilon^{(i)}$ i.i.d. $\mathcal{N}(0, \mathbf{I}_n)$).

Finally, a reminder on chi-squared distribution: if Z_1, \dots, Z_n are independent standard normal $\mathcal{N}(0, 1)$ random variables, then the sum of their squares, $K := \sum_{i=1}^n Z_i^2$, is distributed according to the chi-squared distribution with n degrees of freedom, which is denoted as $K \sim \chi^2(n)$. Also, $\mathbb{E}[K] = n$ and $\text{Var}(K) = 2n$.

1. Generalization bound for a fixed θ using Chebyshev's inequality:

- (a) Show that $Y_j^{(i)} \sim \mathcal{N}(\Phi_{j,:}\theta, 1)$ with $i \in \llbracket m \rrbracket$ and $j \in \llbracket n \rrbracket$.
- (b) Show that $\|Y^{(i)} - \Phi\theta\|_2^2 = (Y^{(i)} - \Phi\theta)^\top (Y^{(i)} - \Phi\theta)$ is distributed as $\chi^2(n)$.
(Hint: rewrite the previous quantity as a sum.)
- (c) Compute the variance of $\|Y^{(i)} - \Phi\theta\|_2^2$.
- (d) Compute the variance of $\widehat{\mathcal{R}}(\theta)$.
(Hint: the elements of the sum $\widehat{\mathcal{R}}(\theta)$ are independent.)
- (e) Using Chebyshev's inequality with $t > 0$, show that:

$$\mathbb{P}\left(\left|\widehat{\mathcal{R}}(\theta) - \mathcal{R}(\theta)\right| > t\right) \leq \frac{\text{Var}\left(\widehat{\mathcal{R}}(\theta)\right)}{t^2} = \frac{2}{mnt^2}.$$

- (f) Solve in $t > 0$, $\delta = \frac{2}{mnt^2}$.
- (g) Rewrite the previous bound as follows:

$$\mathbb{P}\left(\left|\widehat{\mathcal{R}}(\theta) - \mathcal{R}(\theta)\right| \leq \sqrt{\frac{2}{mn\delta}}\right) \geq 1 - \delta.$$

- (h) (Numerical application) set $\delta = 0.01$ in the previous bound and rewrite it as:

$$\mathbb{P}\left(\left|\widehat{\mathcal{R}}(\theta) - \mathcal{R}(\theta)\right| \leq 10\sqrt{\frac{2}{mn}}\right) \geq 0.99.$$